

**Statistical tables are attached**

Two Hours

UNIVERSITY OF MANCHESTER

19 January 2015

14:00 – 16:00

Medical Statistics

MATH38071

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Electronic calculators may be used provided that they cannot store text

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Answer **ALL** five questions in **SECTION A** (40 marks in total).

Answer **TWO** of the three questions in Section B (40 marks in total). If more than two questions from Section B are attempted, then credit will be given for the two best answers.

The total number of marks on the paper is 80.

**A1**

- (i) In studies investigating the effect of an exposure on health, what is the difference between observational and experimental studies? [2 marks]
- (ii) What is a confounding variable? [2 marks]
- (iii) Consider an epidemiological study investigating whether high fat consumption causes heart disease. For such a study suggest an example of each of the three types of variable
  - a) Exposure
  - b) Outcome
  - c) Confounding.

[3 marks]

[Total 7 marks]

**A2.**

A randomized controlled trial is carried out to compare a new treatment regime (*N*) with the existing standard treatment (*S*) for patients. The effectiveness of treatment is assessed by whether the patient is still *infectious after 2 weeks*. The results are summarized in the frequency table below.

		<i>Treatment</i>	
		<i>Standard (S)</i>	<i>New (N)</i>
<i>Infectious after 2 weeks</i>	<i>Yes</i>	100	70
	<i>No</i>	400	430
<i>Total</i>		500	500

- (i) Calculate the rate ratio (RR) of the patient still being *infectious after 2 weeks* with the new treatment (*N*) compared to the standard treatment (*S*). [2 marks]
- (ii) Suppose  $n_T$  patients are randomized to treatment (*T*) and  $n_C$  to the control (*C*). Suppose that the number of events in each of the two treatment groups are  $r_T$  and  $r_C$ . The standard error of

$$\text{the } \log_e [\hat{RR}] \text{ is } \hat{SE} [\log_e [\hat{RR}]] = \sqrt{\frac{1}{r_T} - \frac{1}{n_T} + \frac{1}{r_C} - \frac{1}{n_C}} .$$

Calculate the 95% confidence interval for rate ratio of being *infectious after 2 weeks*.

[5 marks]

- (iii) Is there evidence that the new treatment (*N*) is better than the standard treatment (*S*)?

[2 marks]

[Total 9 marks]

**A3.**

- (i) A researcher designing a randomised control trial considers patient's age and gender to be prognostic. Explain how you would carry out stratified randomisation with stratification by age and sex. [5 marks]
- (ii) What are the advantages and disadvantages of stratified randomisation? [3 marks]
- [Total 8 marks]

**A4.**

The results for a randomised controlled trial comparing a *New* treatment with a *Control* treatment for binary outcome measure (*Recovered after 4 weeks*) are tabulated below. Some patients randomised to the *New* treatment receive the *Control* treatment, but no patients randomised to the *Control* treatment receive the *New* treatment.

<i>Recovered after 4 weeks</i>	<i>Randomised group</i>		
	<i>Received New</i>	<i>Received Control</i>	<i>Control</i>
<i>Yes</i>	110	5	120
<i>No</i>	40	45	80
<i>Total</i>	150	50	200

- (i) Calculate the point estimates of the treatment effect of the *New* treatment compared to the *Control* treatment define by the difference in the proportion of patients who have *Recovered after 4 weeks* for
- an *Intention-To-Treat* analysis
  - a *Per-Protocol* analysis.
  - an *As Treated* analysis
- [4 marks]
- (ii) Drawing on the example above, explain why an *Intention-To-Treat* analysis is preferable to *Per-protocol* and *As-treated* analyses in a superiority trial.
- [4 marks]
- [Total 8 marks]

**A5.**

In a meta-analysis of  $k$  trials, suppose  $\hat{\theta}_i$  is an estimate of the treatment effect for the  $i^{\text{th}}$  study and let

$Var[\hat{\theta}_i]$  be its sampling variance. The minimum variance estimate is defined by  $\hat{\theta}_{MV} = \frac{\sum_i^k w_i \hat{\theta}_i}{\sum_i^k w_i}$ ,

where  $w_i = 1/Var[\hat{\theta}_i]$ , and  $Var[\hat{\theta}_{MV}] = \frac{1}{\sum_{i=1}^k \frac{1}{Var[\hat{\theta}_i]}}$ .

The table below summarizes the outcome of two trials comparing a new drug to prevent high blood cholesterol with the current standard drug treatment. The treatment effect for each study  $(\hat{\theta}_i, i = 1, 2)$  is the difference in mean cholesterol for the two treatments.

- (i) Compute the minimum variance estimate of the overall treatment effect,  $\hat{\theta}_{MV}$ , and determine its 95% confidence interval stating any assumptions you make.

Study (Date of Publication)	Reduction in cholesterol $\hat{\theta}_i$	$Var[\hat{\theta}_i]$
Rahman (2001)	0.36	0.2916
Chung (2008)	0.68	0.1156

[6 marks]

- (ii) What can you conclude from the meta-analysis regarding the performance of the new drug?

[2 marks]

[Total 8 marks]

**B1.**

- (i) For a binary measure
- $Y$
- let
- $\pi$
- and
- $p$
- be the population and sample proportions respectively.

Suppose  $\gamma = \arcsin(\sqrt{\pi})$  and  $\hat{\gamma} = \arcsin(\sqrt{p})$ . Given that  $\frac{d\gamma}{d\pi} = \frac{1}{2\sqrt{\pi(1-\pi)}}$ , use the

approximation  $Var[f(x)] \cong f'(x)^2 \Big|_{x=E[x]} Var[x]$ , to show that  $Var[\hat{\gamma}] \cong \frac{1}{4n}$  where  $n$  is the

sample size.

[6 marks]

- (ii) Consider a parallel group trial with two treatment groups of size
- $n_T$
- and
- $n_C$
- , and a binary outcome measure. Suppose
- $\pi_T, \pi_C, p_T$
- and
- $p_C$
- are the population and sample proportions for each treatment. With the treatment effect defined by
- $\hat{\tau} = \arcsin(\sqrt{p_T}) - \arcsin(\sqrt{p_C})$
- , show

that  $SE[\hat{\tau}] = \sqrt{\frac{1}{4n_T} + \frac{1}{4n_C}}$ . [4 marks]

- (iii) For a normally distributed test statistic
- $T = \frac{\hat{\tau}}{SE[\hat{\tau}]}$
- the power to detect a difference
- $\tau_D$
- with

a two-sided  $\alpha$  size test is given by the expression  $Power = 1 - \Phi\left(z_{\alpha/2} - \frac{\tau_D}{SE[\hat{\tau}]}\right)$ . For a test

of  $H_0 : \pi_T = \pi_C$  vs  $H_1 : \pi_T \neq \pi_C$ , show that the power to detect a difference between two

proportions  $\pi_T$  and  $\pi_C$  can be estimated by  $1 - \Phi\left(z_{\alpha/2} - \frac{\arcsin(\sqrt{\pi_T}) - \arcsin(\sqrt{\pi_C})}{\sqrt{\frac{1}{4n_T} + \frac{1}{4n_C}}}\right)$ ,

stating any required assumptions.

[4 marks]

- (iv) Suppose it is required to design a trial with power
- $(1-\beta)$
- to detect a difference between
- $\pi_T$
- and
- $\pi_C$
- with a two-sided
- $\alpha$
- level test. Show that with two equal size groups, the sample size in each group must be

$$n = \frac{(z_{\alpha/2} + z_{\beta})^2}{2(\arcsin(\sqrt{\pi_T}) - \arcsin(\sqrt{\pi_C}))^2}$$

to achieve the desired power.

[6 marks]

[Total 20 marks]

**B2.**

In a parallel group *non-inferiority* trial a new treatment  $T$  is being compared with a control treatment  $C$  using a normally distributed outcome measure  $Y$ . Assume that large values of  $Y$  represent a worse outcome for the patient. Let  $\mu_T, \bar{y}_T, \mu_C, \bar{y}_C$ , and  $n_C$  be the population means, the sample mean and the sample size for each treatment. Suppose  $\sigma$  is the population standard deviation of both treatments. The treatment effect is defined as  $\tau = \mu_T - \mu_C$ .

- (i) Explain why a significance test of the hypothesis  $H_0 : \tau = 0$  vs  $H_1 : \tau > 0$  would not be appropriate in a *non-inferiority* trial. [3 marks]
- (ii) Suppose that the null hypothesis  $H_0 : \mu_T - \mu_C \geq \tau_N$  is rejected if the upper  $(1-\alpha)$  single sided confidence interval for  $\hat{\tau} (= \bar{y}_T - \bar{y}_C)$  is less than the limit of non-inferiority  $\tau_N$ . Stating

any assumptions, show that 
$$\Pr[\text{Reject } H_0 | \tau] = \Phi\left(\left(\frac{\tau_N}{\sigma\lambda} - z_\alpha\right) - \frac{\tau}{\sigma\lambda}\right),$$

where  $\Phi$  is the cumulative distribution function of the standard normal distribution and 
$$\lambda = \sqrt{1/n_T + 1/n_C}.$$
 [5 marks]

- (iii) Show that  $\Pr[\text{Reject } H_0 | \tau]$  has a maximum under  $H_0$  when  $\tau = \tau_N$ . Hence, show that this procedure has a type I error less than or equal to  $\alpha$ . [7 marks]
- (iv) A randomised controlled non-inferiority trial is carried out to test whether a *New* drug is as effective as a current *Standard* drug for controlling pain. Outcome is measured on a continuous scale with high scores representing greater pain. Eighty patients are randomised to the *New* treatment and 78 to the *Standard* treatment. The statistical output is given below.

Two-sample t test with equal variances

	Obs	Mean	Std. Err.	Std. Dev.	[95% Conf. Interval]	
New	80	45.9	2.62738	23.5	40.67033	51.12967
Standard	78	46.1	2.581592	22.8	40.95939	51.24061
diff		-.2	3.68486		-7.478658	7.078658
diff = mean(New) - mean(Standard)					t = -0.0543	
Ho: diff = 0				degrees of freedom = 156		
Ha: diff < 0		Ha: diff != 0		Ha: diff > 0		
Pr(T < t) = 0.4784		Pr( T  >  t ) = 0.9568		Pr(T > t) = 0.5216		

A difference of 5 mm was considered by researchers to be clinically important difference between treatments. Use the procedure described in part (ii) with a 5% significance level to test whether the *New* drug is non-inferior to the *Standard* drug. [5 marks]

[Total 20 marks]

**B3.**

For an AB/BA crossover trial a model for a continuous outcome  $y_{ij}$  of the  $i^{\text{th}}$  patient in the  $j^{\text{th}}$  period can be written as

$$y_{i1} = \mu + \tau + \xi_i + \varepsilon_{i1} \quad \text{for a patient in sequence AB in period 1,}$$

$$y_{i2} = \mu + \phi + \gamma + \xi_i + \varepsilon_{i2} \quad \text{for a patient in sequence AB in period 2,}$$

$$y_{i1} = \mu + \xi_i + \varepsilon_{i1} \quad \text{for a patient in sequence BA in period 1,}$$

$$y_{i2} = \mu + \tau + \phi + \xi_i + \varepsilon_{i2} \quad \text{for a patient in sequence BA in period 2,}$$

where  $\mu$  is the mean for the sequence BA in period 1,  $\tau$  is the treatment effect of A relative to B,  $\phi$  is the effect of the second period relative to the first,  $\gamma$  is the carryover effect,  $\xi_i$  is a random variable representing patient  $i$  with mean zero and variance  $\sigma_B^2$ , and  $\varepsilon_{ij}$  is the error term for patient  $i$  in period  $j$  assumed to be normally distributed with mean zero and variance  $\sigma_\varepsilon^2$ . Let  $d_i = y_{i2} - y_{i1}$  and let  $\bar{d}_{AB}$ ,  $\mu_{AB}$ ,  $\bar{d}_{BA}$  and  $\mu_{BA}$  be the sample and population means of these for sequences AB and BA respectively.

- (i) Explain what is meant by the term *carryover* effect. [2 marks]
- (ii) In a crossover trial the treatment effect  $\tau$  is estimated by  $\hat{\tau} = (\bar{d}_{BA} - \bar{d}_{AB})/2$ . Show that this will be biased if there is a carryover effect. [4 marks]
- (iii) Let  $a_i = y_{i2} + y_{i1}$  and  $\bar{a}_{AB}$ ,  $\mu_{AB}^A$ ,  $\bar{a}_{BA}$  and  $\mu_{BA}^A$  be the sample and population means for sequences AB and BA respectively. Show that  $E[\bar{a}_{AB} - \bar{a}_{BA}] = \gamma$ . [4 marks]
- (iv) Show that  $Var[\bar{a}_{BA} - \bar{a}_{AB}] = (4\sigma_B^2 + 2\sigma_\varepsilon^2) \left( \frac{1}{n_{BA}} + \frac{1}{n_{AB}} \right)$  [4 marks]
- (v) The test statistic  $T_a$ , defined as  $T_a = \frac{\bar{a}_{AB} - \bar{a}_{BA}}{\hat{SE}[\bar{a}_{AB} - \bar{a}_{BA}]}$ , has been suggested as a test of the hypothesis  $H_0: \gamma = 0$  vs  $H_1: \gamma \neq 0$ . What is the limitation of this as a test for carryover effect? [2 marks]
- (vi) What are the implications of parts (ii) and (iv) for the use of the crossover trial design? [2 marks]
- (vii) Give ONE possible means of preventing a carryover effect in a randomised controlled crossover trial. [2 marks]

[Total 20 marks]

END OF EXAMINATION PAPER