

A.1:

(i) (b) is right.

(ii) We say that $E[X|g]$ is the conditional expectation of X given g if

① $E[X|g]$ is g -measurable (determined),

~~② $E[X|g]$ is integrable, i.e.~~

~~$E[E[X|g]] = E[X]$~~

③ for any event $A \in g$,

$$\int_A E[X|g] dP = \int_A X dP$$

Apply the property ③ to get

$$E[E[X|g]] = \int_{\Omega} E[X|g] dP$$

$$= \int_{\Omega} X dP = E[X] \quad \text{as } \Omega \in g.$$

(!) We say that $\{Z_n, n \geq 0\}$ is a martingale w.r.t.

$\{\mathcal{F}_n, n \geq 0\}$ if

- ① Z_n is \mathcal{F}_n -determined (measurable),
- ② Z_n is integrable, i.e. $E[|Z_n|] < \infty$,
- ③ $E[Z_{n+1} | \mathcal{F}_n] = Z_n$.

(ii) (a)
$$E[Z_{n+3} | \mathcal{F}_n] = E[E[Z_{n+3} | \mathcal{F}_{n+2}] | \mathcal{F}_n]$$

$$= E[E[Z_{n+2} | \mathcal{F}_n] = E[E[Z_{n+2} | \mathcal{F}_{n+1}] | \mathcal{F}_n]]$$

$$= E[Z_{n+1} | \mathcal{F}_n] = Z_n.$$

- (b) ① Z_n is \mathcal{F}_n -determined by the assumption.
 ② Z_n is integrable since $Z_n, n \geq 0$, is a martingale w.r.t. $\mathcal{F}_n, n \geq 0$.

③
$$E[E[Z_{n+1} | \mathcal{F}_n] = E[E[Z_{n+1} | \mathcal{F}_n] | \mathcal{F}_n]]$$

$$= E[Z_n | \mathcal{F}_n] = Z_n$$

Hence $\{Z_n, n \geq 0\}$ is a martingale with respect to $\{\mathcal{F}_n, n \geq 0\}$.

(!)

(1) Since Z_n is a function of X_1, X_2, \dots, X_n , Z_n is \mathcal{F}_n -measurable

(2) As $|Z_n| \leq n + S_n^2 \leq n + (|X_1| + |X_2| + \dots + |X_n|)^2 \leq n + n^2$

Z_n is integrable.

(3) $E[Z_{n+1} | \mathcal{F}_n] = E[S_{n+1}^2 - (n+1) | \mathcal{F}_n]$

$= E[(S_n + X_{n+1})^2 - (n+1) | \mathcal{F}_n]$

$= E[S_n^2 + 2S_n X_{n+1} + X_{n+1}^2 - (n+1) | \mathcal{F}_n]$

$= E[Z_n + 2E[S_n X_{n+1} | \mathcal{F}_n] + E[X_{n+1}^2 - 1 | \mathcal{F}_n]]$

$= Z_n + 2E[S_n X_{n+1} | \mathcal{F}_n] + E[X_{n+1}^2 - 1]$

$= Z_n + 2S_n E[X_{n+1}] + (1 + 1 - 1) - 1$

$= Z_n + 0 = Z_n$

(11)

① Because Z_n is a function of X_1, X_2, \dots, X_n ,

Z_n is f_n -determined.

② As $|Z_n| \leq (\frac{g}{p})^n + (\frac{p}{g})^{-n}$,

Z_n is integrable.

③ $E[Z_{n+1} | \mathcal{F}_n] = E[(\frac{g}{p})^{S_{n+1}} | \mathcal{F}_n]$

$= E[(\frac{g}{p})^{S_n} (\frac{g}{p})^{X_{n+1}} | \mathcal{F}_n] = (\frac{g}{p})^{S_n} E[(\frac{g}{p})^{X_{n+1}} | \mathcal{F}_n]$

$= Z_n E[(\frac{g}{p})^{X_{n+1}}]$

$= Z_n \left\{ \frac{g}{p} P(X_{n+1}=1) + (\frac{g}{p})^{-1} P(X_{n+1}=-1) \right\}$

$= Z_n \left\{ \frac{g}{p} \times p + \frac{g}{p} \times g \right\} = Z_n.$

A.4:

(i) ϕ is self-financing if

$\phi(t) \cdot S(t) = \phi(t+1) \cdot S(t), t \geq 1.$

This means the investor does not bring in or consume

any wealth.

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④

for all $t \geq 1$.

$$\phi(t+1) \cdot \tilde{S}(t) = \phi(t+1) \cdot \tilde{S}(t)$$

Continuing the above procedure, we obtain

$$\text{We deduce that } \phi(1) \cdot \tilde{S}(1) = \phi(2) \cdot \tilde{S}(1).$$

$$= 0 + \tilde{V}_\phi(2) + \phi(1) \cdot \tilde{S}(1) - \phi(2) \cdot \tilde{S}(1).$$

$$= \phi(0) \cdot \tilde{S}(0) - \phi(1) \cdot \tilde{S}(0) + \phi(1) \cdot \tilde{S}(1) + \tilde{V}_\phi(2) - \phi(2) \cdot \tilde{S}(1)$$

$$+ \phi(2) \cdot (\tilde{S}(2) - \tilde{S}(1))$$

$$= \phi(0) \cdot \tilde{S}(0) + \phi(1) \cdot (\tilde{S}(1) - \tilde{S}(0))$$

$$\tilde{V}_\phi(2) = \tilde{V}_\phi(0) + \tilde{G}_\phi(2)$$

Taking $t=2$ we have

$$\phi(0) \cdot \tilde{S}(0) = \phi(1) \cdot \tilde{S}(0).$$

This gives

$$= \tilde{V}_\phi(1) + \phi(0) \cdot \tilde{S}(0) - \phi(1) \cdot \tilde{S}(0).$$

$$= \phi(1) \cdot \tilde{S}(1) + \phi(0) \cdot \tilde{S}(0) - \phi(1) \cdot \tilde{S}(0)$$

$$= \phi(0) \cdot \tilde{S}(0) + \phi(1) \cdot (\tilde{S}(1) - \tilde{S}(0))$$

$$\tilde{V}_\phi(1) = \tilde{V}_\phi(0) + \tilde{G}_\phi(1)$$

Let $t=1$ to get

$$(ii). \text{ Suppose } \tilde{V}_\phi(t) = \tilde{V}_\phi(0) + \tilde{G}_\phi(t), \quad t \geq 1.$$

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(1) For any $n \geq 1$, we have

$$\{T \leq n\} = \bigcup_{k=1}^n \{S_k = -a \text{ or } S_k = b\} \in \mathcal{F}_n$$
 This proves that T is a stopping time.

(ii) ① As a function of X_1, X_2, \dots, X_n , Y_n is \mathcal{F}_n -determined.
 ② $|Y_n| \leq |S_n| + n|p-q| \leq n + n|p-q|$ implies that Y_n is integrable.

$$\textcircled{3} E[Y_{n+1} | \mathcal{F}_n] = E[S_{n+1} - (n+1)(p-q) | \mathcal{F}_n]$$

$$= E[S_n - n(p-q) + X_{n+1} - (p-q) | \mathcal{F}_n]$$

$$= S_n - n(p-q) + E[X_{n+1} - (p-q) | \mathcal{F}_n]$$

$$= Y_n + E[X_{n+1} - (p-q)]$$

$$= Y_n + (p-q) - (p-q) = Y_n$$

(iii) By the definition of T , we have

$$|Z_{T \wedge n}| = \left| \left(\frac{q}{p}\right)^{S_{T \wedge n}} \right| \leq \left(\frac{q}{p}\right)^{a+b} + \left(\frac{q}{p}\right)^{a+b}$$

Thus Doob's Optional Theorem can be applied to get

$$1 = \left(\frac{p}{q}\right)^a (1-p) (S_t^+ = b) + \left(\frac{p}{q}\right)^b p (S_t^+ = b)$$

Hence, it follows that

$$= \left(\frac{p}{q}\right)^a (1-p) (S_t^+ = b) + \left(\frac{p}{q}\right)^b p (S_t^+ = b)$$

$$= \left(\frac{p}{q}\right)^{-a} p (S_t^+ = -a) + \left(\frac{p}{q}\right)^b p (S_t^+ = b)$$

$$E[Z_T] = E\left[\left(\frac{p}{q}\right)^{S_T}\right]$$

On the other hand,

(iv) It is known from (iii) that $E[Z_T] = 1$.

$$E[Y_T] = E[Y_0] = 0$$

Hence the Doob's Optional Theorem implies

$$= |X_{n+1} - (p-q)| \leq |X_{n+1}| + |p-q| \leq 1 + |p-q|$$

$$|Y_{n+1} - Y_n| = |S_{n+1} - S_n - (n+1)(p-q) + n(p-q)|$$

For the martingale $Y_n, n \geq 0$, we have

$$E[Z_T] = E[Z_0] = E\left[\left(\frac{p}{q}\right)^0\right] = 1$$

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Solve the above equation to get

$$P(S_T = b) = \frac{1 - \left(\frac{b}{d}\right)^a}{1 - \left(\frac{b}{p}\right)^a}$$

$$P(S_T = -a) = 1 - P(S_T = b)$$

$$= 1 - \frac{1 - \left(\frac{b}{p}\right)^a}{1 - \left(\frac{b}{d}\right)^a} = \frac{\left(\frac{b}{p}\right)^a - \left(\frac{b}{d}\right)^a}{1 - \left(\frac{b}{d}\right)^a}$$

$E[Y_T] = 0$ is equivalent to

$$E[S_T - T(p-g)] = 0$$

Hence $E[T] = \frac{1}{p-g} E[S_T]$

$$= \frac{p-g}{1} \left\{ (-a)P(S_T = -a) + bP(S_T = b) \right\}$$

$$= \frac{p-g}{1} \left\{ (-a) \frac{1 - \left(\frac{b}{p}\right)^a}{1 - \left(\frac{b}{d}\right)^a} + b \frac{\left(\frac{b}{p}\right)^a - \left(\frac{b}{d}\right)^a}{1 - \left(\frac{b}{d}\right)^a} \right\}$$

$$= \frac{a \left(1 - \left(\frac{b}{p}\right)^a\right) + b \left(\left(\frac{b}{p}\right)^a - \left(\frac{b}{d}\right)^a\right)}{\left(1 - \left(\frac{b}{d}\right)^a\right) \left(b \left(\frac{p}{d}\right)^a - a \left(\frac{d}{p}\right)^a\right)}$$

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(i)

- ① As a function of $N(t)$, $V(t)$ is \mathcal{F}_t -measurable.
- ② $V(t)$ is integrable as

$$E[|V(t)|] \leq E[N(t)] + \lambda t = \lambda t + \lambda t = 2\lambda t$$

③ For $s < t$,

$$E[V(t) | \mathcal{F}_s] = E[V(t) - U(s) + U(s) | \mathcal{F}_s]$$

$$= E[V(t) - U(s)] + U(s)$$

$$= E[N(t)] - \lambda t - E[N(s)] + \lambda s + U(s)$$

$$= \lambda t - \lambda t - \lambda s + \lambda s + U(s) = U(s).$$

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(ii)

- ① Since $V(t)$ is a function of $N(t)$, $V(t)$ is \mathcal{F}_t -determined.
- ② $E[|V(t)|] \leq E[V^2(t)] + \lambda t$

$$= E[(N(t) - \lambda t)^2] + \lambda t = \lambda t + \lambda t = 2\lambda t < \infty$$

③ For $s < t$,

$$E[V(t) | \mathcal{F}_s] = E[V^2(t) - \lambda t | \mathcal{F}_s]$$

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$$= E [(U(t) - U(s) + U(s) - \lambda t)^2 - \lambda t \mid \mathcal{F}_s]$$

$$= E [(U(t) - U(s))^2 + 2(U(t) - U(s))(U(s) - \lambda t) + U(s)^2 - \lambda t \mid \mathcal{F}_s]$$

$$= E [(U(t) - U(s))^2 \mid \mathcal{F}_s] + 2 E [(U(t) - U(s))(U(s) - \lambda t) \mid \mathcal{F}_s]$$

$$+ E [U(s)^2 - \lambda s \mid \mathcal{F}_s] - \lambda(t-s)$$

$$= E [(U(t) - U(s))^2 + 2 U(s) E [U(t) - U(s) \mid \mathcal{F}_s]]$$

$$+ U(s)^2 - \lambda s - \lambda(t-s)$$

$$= E [(N(t) - N(s) - \lambda(t-s))^2 + 2 U(s) E [U(t) - U(s)]]$$

$$+ U(s) - \lambda(t-s)$$

$$= \lambda(t-s) + 2 U(s) [E [N(t)] - \lambda t - E [N(s)] + \lambda s]$$

$$+ U(s) - \lambda(t-s)$$

$$= U(s)$$

(iii)

Again as a function of $N(t)$, $W(t)$ is

\mathcal{F}_t -measurable.

As $N(t)$ has the Poisson distribution with parameter

λt , $E [W(t)] < +\infty$.

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(3)

For $s < t$,

$$E[W(t) | \mathcal{F}_s] = E[\exp[-\lambda(N(t) - N(s)) + \lambda t (1 - e^{-\lambda})]] | \mathcal{F}_s]$$

$$= E[\exp[-\lambda(N(t) - N(s)) + \lambda t (1 - e^{-\lambda})]] | \mathcal{F}_s]$$

$$= W(s) E[W(t) W'(s) | \mathcal{F}_s]$$

$$= W(s) E[\exp[-\lambda(N(t) - N(s)) + \lambda t (1 - e^{-\lambda})]] | \mathcal{F}_s]$$

$$= W(s) E[\exp[-\lambda(N(t) - N(s))]] | \mathcal{F}_s] \exp[\lambda t (1 - e^{-\lambda})]$$

$$= W(s) E[\exp[-\lambda(N(t) - N(s))]] \exp[\lambda t (1 - e^{-\lambda})]$$

$$= W(s) \exp[\lambda t (1 - e^{-\lambda})]$$

$$\times \sum_{k=0}^{\infty} e^{-\lambda k} P(N(t) - N(s) = k)$$

$$= W(s) \exp[\lambda t (1 - e^{-\lambda})]$$

$$\times \sum_{k=0}^{\infty} e^{-\lambda k} \frac{k!}{(\lambda t)^k} e^{-\lambda t}$$

$$= W(s) \exp[\lambda t (1 - e^{-\lambda})]$$

$$\times \sum_{k=0}^{\infty} \frac{k!}{(\lambda t)^k} e^{-\lambda t}$$

$$= W(s) \exp[\lambda t (1 - e^{-\lambda})] e^{-\lambda t}$$

$$= W(s) \exp[\lambda t (1 - e^{-\lambda})] e^{-\lambda t}$$

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(11)

$$= W(s)$$

B7.

(i) We say that $\phi = (\phi_0(t), \phi_1(t))$ is an arbitrage opportunity if $V_\phi(0) = 0$, and

$$V_\phi(T) \geq 0 \text{ and } P(V_\phi(T) > 0) > 0.$$

(ii) The market is free of arbitrage if

and only if there exists a probability measure

\mathbb{Q} under which the discounted price \tilde{S}_1 is

a martingale.

Suppose $\mathbb{Q}(\{w_1\}) = p^*$, $\mathbb{Q}(\{w_2\}) = q^* = 1 - p^*$.

In order that \tilde{S}_1 is a martingale under \mathbb{Q} we

must have

$$E^{\mathbb{Q}^*}[\tilde{S}_1] = E^{\mathbb{Q}^*}[S_0] = E^{\mathbb{Q}^*}\left[\frac{S_1^{(1)}}{S_0^{(1)}}\right]$$

$$= \frac{s}{d}$$

$$\text{But } E^{\mathbb{Q}^*}[\tilde{S}_1] = E^{\mathbb{Q}^*}\left[\frac{S_1^{(1)}}{S_0^{(1)}}\right]$$

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$$= \frac{S_1(1, w_1)}{S(1+r)} Q^*(\{w_1\}) + \frac{S_1(1, w_2)}{S(1+r)} Q^*(\{w_2\})$$

$$= \frac{d_1}{S(1+r)} p^* + \frac{d_2}{S(1+r)} (1-p^*)$$

Thus we have the following equation for p^* :

$$\frac{d_1}{S(1+r)} p^* + \frac{d_2}{S(1+r)} (1-p^*) = \frac{d}{S}$$

Solve the above equation to get

$$p^* = \frac{d_2 - d_1}{d_2 - d_1}$$

In order that Q is a well defined martingale probability, we must have

$$0 < p^* = \frac{d_2 - d_1}{d_2 - d_1} < 1$$

This is equivalent to

$$\begin{cases} d_2 - d(1+r) > 0 \\ d_2 - d(1+r) < d_2 - d_1 \end{cases}$$

Combining these two we get

$$d_1 < d(1+r) < d_2$$

The condition is:

$$d_1 < d(1+r) < d_2$$

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The martingale probability measure is given by (14)

$$Q(\omega_1) = p^* = \frac{d_2 - d_1}{d_2 - d_{(1+r)}}$$

$$Q(\omega_2) = 1 - p^* = \frac{d_{(1+r)} - d_1}{d_2 - d_1}$$

(iii) An European call option with strike price K gives

one the right to buy on the expiry date at a specified price K .

The payoff (claim) is

$$X = \begin{cases} S_{(1)} - K & \text{if } S_{(1)} \geq K \\ 0 & \text{if } S_{(1)} < K \end{cases}$$

(iv)

(a) Under the assumption, the martingale probability is

$$Q(\omega_1) = p^* = \frac{8-5}{8-3} = \frac{3}{5}$$

$$Q(\omega_2) = 1 - p^* = \frac{2}{5}$$

The price of the call option is

$$\pi(0) = S_0(0) = E^Q \left[\frac{X}{S_0(1)} \right]$$

$$= 5 E^Q \left[\frac{X}{5(1+r)} \right] = \frac{1}{1+r} E^Q[X]$$

$$= \frac{1}{1+r} \{ X(\omega_1) Q(\omega_1) + X(\omega_2) Q(\omega_2) \}$$

Solve the above equation to obtain $\phi_1 = \frac{1}{5}$, $\phi_0 = -\frac{3}{25(1+r)}$.

i.e.
$$\begin{cases} \phi_0 5(1+r) + \phi_1 d_1 = 0 \\ \phi_0 5(1+r) + \phi_2 d_2 = 8-7=1 \end{cases}$$

$$\begin{cases} \phi_0 5(1+r) + \phi_1 d_1 = X(w_1) = (d_1 - k)^+ = 0 \\ \phi_0 5(1+r) + \phi_2 d_2 = X(w_2) = (d_2 - k)^+ = 1 \end{cases}$$

Let $w = w_1$ to get w_1 and $w = w_2$ respectively

$$V\phi(1) = \phi_0 S_0(1) + \phi_1 S_1(1) = X(w)$$

Then of the X.

(b) Let $\phi = (\phi_0, \phi_1)$ be the replicating strategy for the claim

$$\begin{aligned} &= \frac{1}{1+r} \left\{ \theta (d_1 - k)^+ + \frac{5}{2} (d_2 - k)^+ \right\} \\ &= \frac{1}{1+r} \left\{ (3-7)^+ \times \frac{5}{2} + \frac{5}{2} (8-7)^+ \right\} \\ &= \frac{1}{1+r} \times \frac{5}{2} = \frac{5(1+r)}{2} \end{aligned}$$